87[K].-Minoru Siotani \& Masaru Ozäwa, "Tables for testing the homogeneity of $k$ independent binomial experiments on a certain event based on the range," Ann. Inst. Stat. Math., v. 10, 1958, p. 47-63.
Let $k$ series of $N$ trials each of a certain event be performed with the outcome of $\nu_{i}$ occurrences in the $i$-th series in which the fixed probability of occurrence was $p_{i}, i=1,2, \cdots, k$. To test the null hypothesis of homogeneity:

$$
p_{1}=p_{2}=\cdots=p_{k}=p
$$

Siotani had previously proposed the statistic, $R_{k}(N, p)$, the range of the $\nu_{i}[1]$. The tables in this paper give for $N=10(1) 20,22,25,27,30 ; k=2(1) 15$;

$$
p=.1(.1) .5
$$

$\alpha=.001, .005, .01(.01) .06, .08, .1$, the greatest $r_{k}$ for which

$$
\operatorname{Pr}\left\{R_{k}(N, p) \geqq r_{k}\right\}<\alpha+.0005
$$

The cases in which for the $r_{k}$ given, $\alpha<\operatorname{Pr}\left\{R_{k}(N, p) \geqq r_{k}\right\}<\alpha+.0005$ or

$$
\alpha-.005<\operatorname{Pr}\left\{R_{k}(N, p) \geqq r_{k}\right\}<\alpha
$$

are indicated by attaching $\mathrm{a}+$ or a - respectively to the value of $r_{k}$.

> C. C. Craig

University of Michigan
Ann Arbor, Michigan

1. Minoru Siotani, "Order statistics for discrete case with a numerical application to the binomial distribution," Ann. Inst. Stat. Math., v. 8, 1956, p. 95-104.
$\mathbf{8 8}[\mathrm{K}]$.-P. N. Somerville, "Tables for obtaining non-parametric tolerance limits," Ann. Math. Stat., v. 29, 1958, p. 599-601.
Let $P$ be the fraction of a population having a continuous but unknown distribution function that lies between the $r$-th smallest and the $s$-th largest values in a random sample of $n$ drawn from that population. Then for any $r, s \geqq 0$ such that $r+s=m$, Table I gives the largest value of $m$ such that with confidence coefficient $\geqq \gamma$ we may assert that $100 P \%$ of the population lies in the interval $(r, s)$ for $\gamma=.5, .75, .9, .95, .99$ and $n=50(5) 100(10) 150,170,200(100) 1000$. Table II gives $\gamma$ to 2D for the assertion that $100 \mathrm{P} \%$ of the population lies within the range, $(r, s=1)$, in a sample of $n$ for $P=.5, .75, .9, .95, .99$ and

$$
n=3(1) 20,25,30(10) 100 .
$$

C. C. Craig

University of Michigan
Ann Arbor, Michigan
$\mathbf{8 9}[\mathrm{K}]$.-G. P. Steck, "A table for computing trivariate normal probabilities," Ann. Math. Stat., v. 29, 1958, p. 780-800.
Let $X, Y, Z$ be standardized random variables obeying a trivariate normal dis-
tribution law. The author finds $\operatorname{Pr}(X \leqq h, Y \leqq k, Z \leqq m)$ in terms of three functions:

$$
\begin{aligned}
& G(x)=(2 \pi)^{-1 / 2} \int_{-\infty}^{x} \exp \left(-\frac{x^{2}}{2}\right) d x, \quad T(h, a) \\
&=(2 \pi)^{-1} \int_{0}^{a}\left[\exp \left\{-h^{2}\left(1+x^{2}\right) / 2\right\}\right]\left(1+x^{2}\right)^{-1} d x
\end{aligned}
$$

and

$$
S(h, a, b)=\int_{-\infty}^{h} T(a s, b) G^{\prime}(s) d s
$$

The $T$-function has been tabulated by D. B. Owen [1, 2] and a table of $S(m, a, b)$ is given in the present paper to 7 D for $a=0(.1) 2(.2) 5(.5) 8, b=.1(.1) 1$ and a range of values of $m$ decreasing from $0(.1) 1.5, \infty$ for $a=0(.1) 1.2$ to $0(.1) .3, \infty$ for $a=6(.5) 8$. The tabulated values are believed accurate to 0.6 in the seventh decimal place. There is considerable discussion of the main problem, of properties of and relations among the functions used, and a numerical example is worked out. The method of construction of the table is given and the efficacy of linear interpolation in it is discussed.
C. C. Craig

University of Michigan
An Arbor, Michigan

1. D. B. Owen, The Bivariate Normal Probability Distribution, Office of Technical Services, Department of Commerce, Washington, D. C., 1957, [MTAC Review 134, v. 12, 1958, p. 285-286.]
2. D. B. Owen, "Tables for computing bivariate normal probabilities," Ann. Math. Stat., v. 27, 1956, p. 1075-1090. [MTAC Review 135, v. 12, 1958, p. 286.]
$\mathbf{9 0 [ K}]$.-G. Taguti, "Tables of tolerance coefficients for normal populations,"
Union of Japanese Scientists and Engineers, Reports of Statistical Application
Research, v. 5, 1958, p. 73-118.
The tolerance limits $T_{1}, T_{2}$ are to be determined so that with probability $1-\alpha$ the interval ( $T_{1}, T_{2}$ ) includes a given fraction, $P$, of the population. Following the method of Wald \& Wolfowitz [1] for a sample from $N\left(\mu, \sigma^{2}\right), T_{1}$ and $T_{2}$ are found by $T_{1}=\hat{\mu}-k \sqrt{S_{e} / \nu}$ and $T_{2}=\hat{\mu}+k \sqrt{S_{e} / \nu}$, in which $\hat{\mu}$ is an unbiased estimate of $\mu$ with variance $\sigma^{2} / n$ and $S_{\epsilon}$ is an independent error sum of squares with $\nu$ degrees of freedom. As illustrated by the author this permits useful applications in which $n$ is not simply the sample size and $\nu=n-1$ as is the case for the tables of Bowker [2]. The present tables give $k$ to 3 S for $P=.9, .95, .99,1-\alpha=$ $.9, .95, .99, n=.5(.5) 2(1) 10(2) 20(5) 30(10) 60(20) 100,200,500,1000, \infty$ and $\nu=1(1) 20(2) 30(5) 100(100) 1000, \infty$. The calculations were done with a slide rule and the author fears there may be errors up to one per cent. Some cursory comparisons with Bowker's tables for $\nu=n-1$ showed frequent differences in the third significant figure.
C. C. Craig

University of Michigan
Ann Arbor, Michigan

